

PAC Style Guarantees for Doubly Robust Generalized Front-Door Estimator

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Applications of Causal Inference

■ Healthcare:

- Understanding the effect of treatments and interventions.
- Example: Determining if a new drug reduces the risk of a disease.

■ Public Policy:

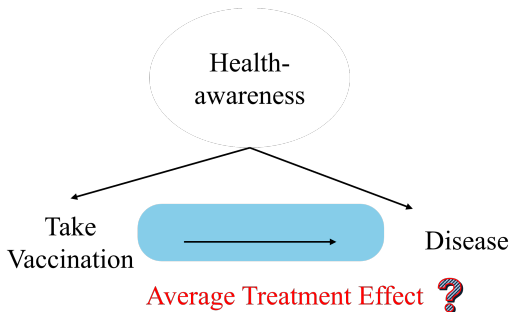
- Evaluating the impact of policies and programs.
- Example: Assessing the effectiveness of a new education policy on student performance.

■ Business:

- Identifying strategies that increase sales or customer retention.
- Example: Measuring the effect of a marketing campaign on product sales.

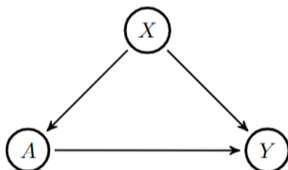
Background on Causal Inference

Example: understand average treatment effect in health care

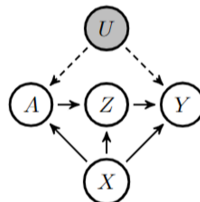


Background: Plausible Causal Models

The causal effect can be identified using backdoor, front door models



(a) G_1 : Back-door



(b) G_2 : Front door

Example: ATE estimation with back-door model

■ Outcome Regression Model

- Predicts the potential outcomes given covariates and treatment.
- Estimates $E[Y|A, X]$.

Background: ATE Estimation

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- Estimates the probability of receiving the treatment given covariates.
- Formally, $e(X) = P(A = 1|X)$.

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■ Propensity Score Model

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However:

- Model misspecification is common \Rightarrow **Incorrect estimation**.
- Slow convergence is common \Rightarrow **Not \sqrt{n} -consistency**

Introduction to Doubly Robust Estimators

- **Recent advance: Doubly robust estimator**
 - Doubly robust estimators combine **inverse probability weighting (IPW)** and **outcome regression** to estimate causal effects;

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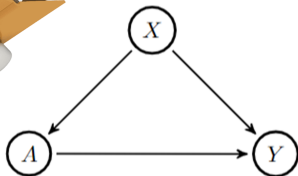
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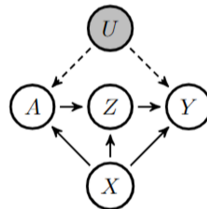
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- Enable us to construct valid **confidence interval** for our treatment effect estimates;
- Introduce a **\sqrt{n} -consistent estimator**
 - As $n \rightarrow \infty$, the estimation error $\hat{\psi} - \psi^*$ goes to zero at a rate of $n^{-1/2}$;
 - We really like our estimators to be at least \sqrt{n} -consistent.

Background: ATE Estimation

Apply doubly robust estimator!!



(a) G_1 : Back-door









(b) G_2 : Front door

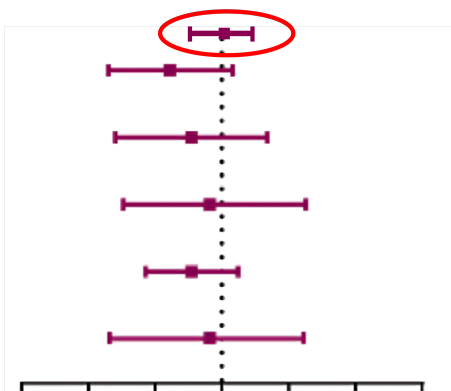
Our Motivation

However, only have finite observations



	Take vaccination	Take placebo	Individual Treatment effect
	15.48	-	-
	13.52	-	-
	-	16.2	-
	-	9.58	-

Average	12.7	10.52	2.18 

Our Motivation



Background: Introduction to PAC Learning

Probably Approximately Correct (PAC) Learning is a framework provides a theoretical foundation for understanding finite sample complexity.

- Accuracy (ϵ): the maximum allowed error.
- Confidence (δ): the probability that the learned hypothesis is approximately correct.
- Sample complexity: the number of samples(n) required to achieve (ϵ, δ) guarantees.

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Goal: With high probability ("Probably"), the selected hypothesis will have lower error ("Approximately Correct")

Our Problem

-
- PAC-style guarantees for doubly machine learning estimator



Back-door Model

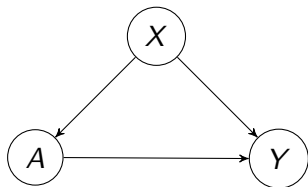


Figure: Back-door adjustment

- \mathcal{A} be $\{0, 1\}$
- \mathcal{X}, \mathcal{Y} be any sets
- X be a random variable on \mathcal{X}
- $A(x), \hat{A}(x)$ be random variables on \mathcal{A} for any $x \in \mathcal{X}$
- $Y(a, x), \hat{Y}(a, x)$ be random variables on \mathcal{Y} for any $a \in \mathcal{A}, x \in \mathcal{X}$

Back-door adjustment: Estimator

The causal effect of A on Y denote as:

$$\psi^* = \mathbb{E}[Y \mid \text{do}(A = a^*)]$$

we define:

$$\phi(a, x, y) = \mathbb{E}_{\hat{Y}(a^*, x)} \hat{Y} + \frac{1[a = a^*]}{\hat{A}(a; x)} \left(y - \mathbb{E}_{\hat{Y}(a, x)} \hat{Y} \right)$$

and

$$\psi_n = \frac{1}{n} \sum_{i=1}^n \phi_{\hat{A}, \hat{Y}}(a_i, x_i, y_i)$$

Back-door adjustment: Goal

We revisit the well-known double/debiased machine learning (DML) estimator for covariate adjustment in the BD setting [RRZ94; Che+17] and analyze the mean-squared error of $|\psi_n - \psi^*|$ in the finite sample setting.

Specifically, given finite samples, we want to bound MSE error in term of the error in \widehat{A}, \widehat{Y} .

Back-door adjustment: Result

Definition

Given two distributions P and Q , and an even integer $p \geq 2$, we define the χ^p divergence between them as:

$$\chi^p(P\|Q) = \mathbb{E}_{x \sim P} \left[\left(1 - \frac{Q(x)}{P(x)} \right)^p \right].$$

Novelty: The novelty of our result is that we express the mean-squared error explicitly in terms of the errors in the estimates of the treatment and outcome distributions. These errors are formulated in terms of χ^2 -divergence.

Back-door adjustment: Result

Assumption

Assume for all x , the following condition holds:

$$\mathbb{E}_{Y(a^*, x)} Y^2 \leq V, \quad \Pr[A(x) = a^*] \geq \mu, \quad \text{and} \quad \Pr[\hat{A}(x) = a^*] \geq \mu.$$

Theorem

Under the above Assumption, for any $\varepsilon > 0$:

$$\Pr[|\psi_n - \psi^*| > \varepsilon] < \frac{1}{n\varepsilon^2} \mathcal{O}_{V, \mu} \left(1 + \mathbb{E}_x \chi^2 \left(\hat{Y}(a^*, x) \| Y(a^*, x) \right) + \mathbb{E}_x \chi^2 \left(\hat{A}(x) \| A(x) \right) \right) \\ + \frac{1}{\varepsilon^2} \mathcal{O}_{V, \mu} \left(\mathbb{E}_x \chi^2 \left(\hat{Y}(a^*, x) \| Y(a^*, x) \right) \cdot \chi^2 \left(\hat{A}(x) \| A(x) \right) \right)$$

Back-door adjustment: Result

What does this result mean?

$$\begin{aligned} & \Pr[|\psi_n - \psi^*| > \varepsilon] \\ & < \frac{1}{n\varepsilon^2} \mathcal{O}_{V,\mu} \left(\underbrace{1}_{\text{Part1}} + \underbrace{\mathbb{E}_x \chi^2 \left(\widehat{Y}(a^*, x) \| Y(a^*, x) \right)}_{\text{Part2}} + \underbrace{\mathbb{E}_x \chi^2 \left(\widehat{A}(x) \| A(x) \right)}_{\text{Part3}} \right) \\ & \quad + \underbrace{\frac{1}{\varepsilon^2} \mathcal{O}_{V,\mu} \left(\mathbb{E}_x \chi^2 \left(\widehat{Y}(a^*, x) \| Y(a^*, x) \right) \cdot \chi^2 \left(\widehat{A}(x) \| A(x) \right) \right)}_{\text{Part4}} \end{aligned}$$

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- Part 1: error incurred by the oracle estimator;

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- Part 1: error incurred by the oracle estimator;
- Part 2 & Part 3: Mismatch measured in χ^2 between the model estimates and the truth, for the outcome and propensity distributions;

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- Part 1: error incurred by the oracle estimator;
- Part 2 & Part 3: Mismatch measured in χ^2 between the model estimates and the truth, for the outcome and propensity distributions;
- Part 4: *Mixed-bias* or *product rate* phenomenon of the doubly robust estimators [Che+20]

Front-door Criterion - Model

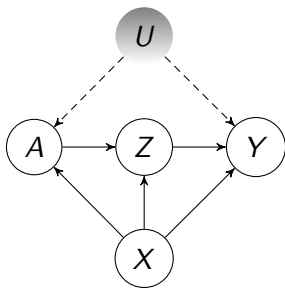


Figure: Front-door adjustment

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- $Y(z, a, x), \hat{Y}(z, a, x)$ be random variables on \mathcal{Y} for any $a \in \mathcal{A}, z \in \mathcal{Z}, x \in \mathcal{X}$

Front-door adjustment - Notations

- we use ψ^* denote the ground truth estimator
- $\phi_{Y,Z,A}$ to denote the oracle estimator
- $\phi_{\hat{Y},\hat{Z},\hat{A}}$ denote the empirical estimator.
- $\psi^* = \mathbb{E}[Y \mid \text{do}(A = a^*)]$
- $\psi_n = \frac{1}{n} \sum_{i=1}^n \phi(a_i, z_i, y_i, x_i)$

Our goal is to bound the MSE error of $(\psi_n - \psi^*)$ in terms of the error in $\hat{A}, \hat{Z}, \hat{Y}$.

Front-door adjustment- Estimators

ψ^* can be identified in terms of the observable distribution in a number of ways, all equivalent to each other.

- The first, ψ_1 , doesn't require $A | X$:

$$\psi_1 = \mathbb{E}_{A,X} \left[\frac{\mathbb{E} \left[\mathbb{E}[Y | A, Z, X] | a^*, X \right]}{Z} \right].$$

- The second, ψ_2 , doesn't require $Z | A, X$:

$$\psi_2 = \mathbb{E}_{A,Z,X} \left[\frac{1[A = a^*]}{\Pr[A | X]} \cdot \mathbb{E}_{A'} \left[\frac{\mathbb{E}[Y | A', Z, X] | X}{Y} \right] \right].$$

- The third, ψ_3 , doesn't require $Y | A, Z, X$:

$$\psi_3 = \mathbb{E}_{Y,A,Z,X} \left[Y \cdot \frac{\Pr[Z | a^*, X]}{\Pr[Z | A, X]} \right]$$

Front-door adjustment - DR estimator

Define the following quantity [Ful+20]:

$$\begin{aligned} & \phi(a, z, y, x) \\ &= \mathbb{E}_{\widehat{Z}(a^*, x)} \left[\mathbb{E}_{\widehat{Y}(a, \widehat{Z}, x)} \widehat{Y} \right] + \frac{1[a = a^*]}{\widehat{A}(a; x)} \mathbb{E}_{\widehat{A}(x)} \left(\mathbb{E}_{\widehat{Y}(\widehat{A}, z, x)} \widehat{Y} - \mathbb{E}_{\widehat{Z}(a, x)} \left[\mathbb{E}_{\widehat{Y}(\widehat{A}, \widehat{Z}, x)} \widehat{Y} \right] \right) \\ & \quad + \left(y - \mathbb{E}_{\widehat{Y}(a, z, x)} \widehat{Y} \right) \cdot \frac{\widehat{Z}(z; a^*, x)}{\widehat{Z}(z; a, x)} \end{aligned}$$

We can get a "doubly-robustness" type property for ϕ : If any two of \widehat{A} , \widehat{Z} , \widehat{Y} are always correct, then $\mathbb{E}_{A, Z, Y, X} \phi = \psi^*$.

Front-door adjustment - Our Result

Assumption

Assume $\forall a, z, x,$

$$\mathbb{E}_{Y(a,z,x)} Y^2 \leq V, \Pr[Z(a,x) = z] \geq \mu_Z \cdot 1[\Pr[Z(a^*,x) = z] \neq 0],$$

and $\Pr[A(x) = a^*] \geq \mu_A$

Front-door adjustment - Our results

Theorem

$$\begin{aligned} & \Pr[|\psi_n - \psi^*| > \varepsilon] \\ & < \frac{1}{n\varepsilon^2} \mathcal{O}_{V, \mu_Z, \mu_A} \left(1 + \mathbb{E}_{a, z, x} \chi^2 \left(\widehat{Y}(a, z, x) \| Y(a, z, x) \right) + \mathbb{E}_x \chi^2 \left(\widehat{A}(x) \| A(x) \right) \right. \\ & \quad \left. + \mathbb{E}_x \chi^2 \left(\widehat{Z}(a^*, x) \| Z(a^*, x) \right) + \mathbb{E}_{a, x} \chi^2 \left(\widehat{Z}(a, x) \| Z(a, x) \right) \right) \\ & \quad + \frac{1}{\varepsilon^2} \mathcal{O}_{V, \mu_A, \mu_Z} \left(\left(\sqrt{\mathbb{E}_{a, z, x} \chi^4 \left(\widehat{Y}(a, z, x) \| Y(a, z, x) \right)} + \sqrt{\mathbb{E}_x \chi^4 \left(\widehat{A}(x) \| A(x) \right)} \right. \right. \\ & \quad \left. \left. + \sqrt{\mathbb{E}_x \chi^4 \left(\widehat{Z}(a^*, x) \| Z(a^*, x) \right)} + \sqrt{\mathbb{E}_{a, x} \chi^4 \left(\widehat{Z}(a, x) \| Z(a, x) \right)} \right)^2 \right) \end{aligned}$$

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What does this result mean?

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- Learning a distribution that minimizes a particular divergence is a question in *distribution learning*
- For example, the problem of learning a distribution minimizing the χ^2 divergence was explicitly studied in [Kam+15].

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- Extend the current work to more general graph?
- How to analyze the mixed-bias term in a "simpler way"?
- Real-data experiments.

Thank You!

Q & A

References I

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Backup

Back-door adjustment

$$\begin{aligned}\Delta_1 &= \frac{1}{n} \sum_i \left(\mathbb{E}_{\hat{Y}(a^*, x_i)} \hat{Y} - \mathbb{E}_{Y(a^*, x_i)} Y - \frac{1[a_i = a^*]}{\mathcal{A}(a_i; x_i)} \left(\mathbb{E}_{\hat{Y}(a_i, x_i)} \hat{Y} - \mathbb{E}_{Y(a_i, x_i)} Y \right) \right) \\ \Delta_2 &= \frac{1}{n} \sum_i 1[a_i = a^*] \left(\frac{1}{\hat{\mathcal{A}}(a_i; x_i)} - \frac{1}{\mathcal{A}(a_i; x_i)} \right) \cdot \left(y_i - \mathbb{E}_{Y(a_i, x_i)} Y \right) \\ \Delta_3 &= -\frac{1}{n} \sum_i \left(\mathbb{E}_{\hat{Y}(a_i, x_i)} \hat{Y} - \mathbb{E}_{Y(a_i, x_i)} Y \right) \left(\frac{1}{\hat{\mathcal{A}}(a_i; x_i)} - \frac{1}{\mathcal{A}(a_i; x_i)} \right)\end{aligned}$$

Back-door adjustment

We bound each of $\Delta_1, \Delta_2, \Delta_3$ by explicit computation.

$$\mathbb{E}[\Delta_1^2] \leq \mathcal{O}_{V,\mu} \left(\frac{1}{n} \cdot \mathbb{E}_x \chi^2 \left(\widehat{Y}(a^*, x) \| Y(a^*, x) \right) \right)$$

$$\mathbb{E}[\Delta_2^2] \leq \mathcal{O}_{V,\mu} \left(\frac{1}{n} \cdot \mathbb{E}_x \chi^2 \left(\widehat{A}(x) \| A(x) \right) \right)$$

$$\mathbb{E}[\Delta_3^2] \leq \mathcal{O}_{V,\mu} \left(\mathbb{E}_x \chi^2 \left(\widehat{Y}(a^*, x) \| Y(a^*, x) \right) \cdot \mathbb{E}_x \chi^2 \left(\widehat{A}(x) \| A(x) \right) \right)$$

